

# Weak Identification Robust Methods for Production Function Estimation

Jorge de la Cal Medina

University of Amsterdam, Tinbergen Institute

August 27, 2025

- ▶ **Production function (PF) estimation:**

$$y = x'\beta + \omega + \varepsilon$$

Firms choose inputs  $x$  after observing state  $\omega$  (Marschak and Andrews, 1944).

- ▶ **Control function approach (CFA):** Proxying  $\omega$  with a control function allows to estimate  $\beta$  with GMM (Akerberg et al., 2015).
- ▶ **Weak identification:** When proxies have low explanatory power (measurement error) Jacobian near-singular, leading to biased, non-normal estimates and invalid standard inference.

## Contributions

1. Identification with noisy proxy feasible (under restrictions)
2. Weak proxies (small signal/noise) can cause biased estimates with non-normal distributions.
3. Adapt weak identification-robust methods: bootstrap tests for proxy strength (Angelini et al., 2024) and robust confidence sets (Stock and Wright, 2000).

## Related Literature

### Production-function identification

- ▶ Olley and Pakes (1996) (OP), Levinsohn and Petrin (2003) (LP), Akerberg et al. (2015) (ACF) discuss identification using (strong) proxies; Gandhi et al. (2020) (GNR) discuss non-identification (in nonparametric models)
- ▶ *We consider sources for weak identification*

### Weak identification in structural models

- ▶ DSGE (Canova & Sala 2009), NK Phillips curve (Mavroeidis, Plagborg-Møller & Stock 2013), BLP demand (Armstrong 2016)
- ▶ *We extend this list with structural production-function estimation*

### Empirical work

- ▶ Empirical applications using control function estimator
  - ▶ Trade: Pavcnik (2002); Topalova (2010); Fernandes (2007)
  - ▶ Market power: De Loecker et al. (2020); Autor et al. (2020)
  - ▶ Finance: Gopinath et al. (2017); Gourinchas et al. (2020)
- ▶ *We give guidance on model specification and add to toolbox*

# Structure of Presentation

- ▶ Identification Analysis
  - ▶ Present identification strategy of CFA
  - ▶ Weak proxies
- ▶ Monte Carlo simulation
- ▶ Empirical Analysis
- ▶ Conclusion

# Outline

Introduction

**Identification Analysis**

Monte Carlo

Empirical application

Conclusion

# The Control Function Approach

- ▶ General PF specification for firm  $i$  at period  $t$ :

$$y_{it} = x'_{it}\beta + \omega_{it} + \varepsilon_{it}, \quad \mathbb{E}[\varepsilon_{it} | \mathcal{I}_{it}] = 0 \quad (1)$$

- ▶ Markov process for state variable

$$\omega_{it} = \rho\omega_{it-1} + \xi_{it}, \quad \mathbb{E}[\xi_{it} | \mathcal{I}_{it-1}] = 0 \quad (2)$$

- ▶ Demand function  $h$  of intermediate inputs  $m$

$$m_{it} = h(\omega_{it}, x_{it}) + \nu_{it} \quad (3)$$

- ▶ strictly monotonic in  $\omega_{it}$
- ▶ allow for unobservable  $\nu_{it}$  (unlike OP, LP, ACF)

- ▶ Control function

$$\begin{aligned} \omega_{it} &= h^{-1}(m_{it} - \nu_{it}, x_{it}) \\ &= \Psi(m_{it}, x_{it}) + \eta_{it}, \quad \mathbb{E}[\eta_{it} | x_{it}, m_{it}] = 0 \end{aligned}$$

- ▶  $\Psi(m, x) = E[\omega_{it} | m_{it} = m, x_{it} = x]$

# Estimation

- ▶ Estimation equation

$$y_{it} = x'_{it}\beta + \rho(y_{it-1} - x'_{it-1}\beta) - \rho\varepsilon_{it-1} + \xi_{it} + \varepsilon_{it} \quad (4)$$

- ▶ GMM Moment conditions

$$\mathbb{E}\left[z_{it}(y_{it} - x'_{it}\beta - \rho(y_{it-1} - x'_{it-1}\beta))\right] = 0$$

- ▶ Instrument vector  $z_{it}$  contains  $m_{it-1}$  to instrument for  $y_{it-1}$  (through control function)
- ▶ Straightforward to get CUE (Hansen et al., 1996)

# Nonidentification & Weak Proxies

## Identification

- ▶ requires full-rank Jacobian

$$\mathcal{J} := \left( \mathbb{E}[z_{it}(x'_{it} - \rho_0 x'_{it-1})] \vdots \mathbb{E}[z_{it}(y_{it-1} - x'_{it-1}\beta_0)] \right)$$

## Nonidentification

- ▶ Assume linear control function  $\Psi(m, x) = \gamma_m m + x' \gamma_x$
- ▶  $\mathcal{J}$  is rank-deficient, if
  - ▶  $x_{it}$  and  $x_{it-1}$  collinear
  - ▶  $\gamma_m \rightarrow 0$

## Weak proxies

- ▶ Small  $\gamma_m$  (weak proxies) lead to near rank-deficiency  $\Rightarrow$  Weak identification
- ▶ Example: classical measurement error in proxies attenuates  $\gamma_m$

# Outline

Introduction

Identification Analysis

Monte Carlo

Empirical application

Conclusion

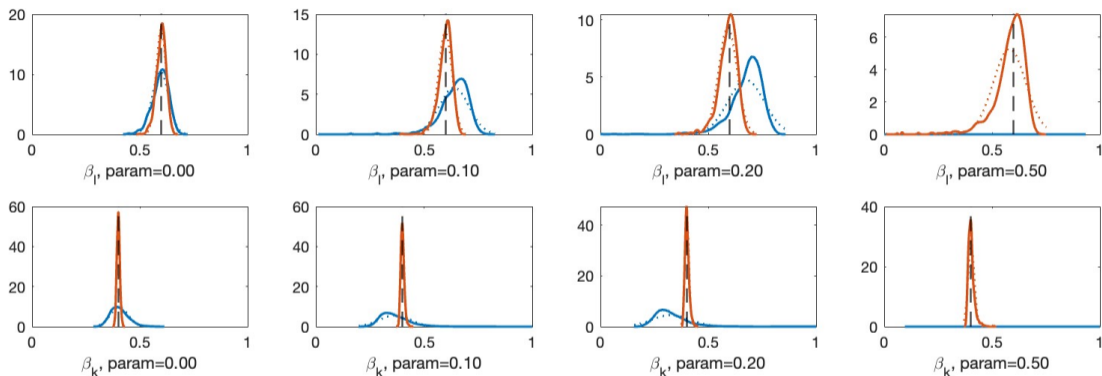
## Monte Carlo

- ▶ Dynamic model of firm investment and production (ACF)
  - ▶ Two inputs capital  $k$  and labor  $l$
  - ▶ Intermediate inputs,  $m$ , subject to measurement error
  - ▶ Tweak: reduce optimization error in  $l$  from 10% to 5%
- ▶ Estimators
  - ▶ ACF estimator
  - ▶ CUE

**Table 1:** Bias of simulated estimates (with standard errors). Measurement error in terms of additional % variance of  $m$ .

Meas.	ACF		CUE	
	$\beta_l$	$\beta_k$	$\beta_l$	$\beta_k$
0.0	-0.005 (0.038)	+0.005 (0.041)	-0.001 (0.023)	0.000 (0.007)
0.1	-0.500 (0.300)	-0.301 (0.299)	+0.038 (0.068)	-0.035 (0.070)
0.2	-0.445 (0.985)	-0.493 (0.983)	-0.461 (0.346)	-0.300 (0.300)
0.5	—	—	-0.532 (1.014)	-0.436 (0.968)

# Monte Carlo Estimates of PF Parameters



1000 replications. ACF in blue. CUE in red. Empirical densities: solid. Normal pdf: dotted. Panels ordered left to right for larger measurement-error magnitudes.

► Measurement error  $\uparrow$ : stronger bias and non-normality

# Monte Carlo: Inference on PF Parameters

Table 2: Rejection frequencies at 5% nominal level

Meas.	Joint test (non-robust) <sup>1</sup>	Joint test (robust) <sup>2</sup>	BS joint normality <sup>3</sup>
0.0	0.064	0.060	0.057
0.1	0.059	0.048	0.068
0.2	0.075	0.062	0.167
0.5	0.128	0.050	0.431

Based on 1000 replications. <sup>1</sup>Wald. <sup>2</sup>Stock and Wright (2000). <sup>3</sup>Doornik and Hansen (2008).

- ▶ Measurement error ↑:
  - ▶ Non-robust test: size  $> 5\%$
  - ▶ Robust test: size  $\approx 5\%$
  - ▶ BS normality test: rejections ↑

# Outline

Introduction

Identification Analysis

Monte Carlo

**Empirical application**

Conclusion

## Data & empirical setting (Chile & U.S.)

- ▶ Replicate Raval (2023) and apply CUE estimator and identification-robust methods
- ▶ Chile (ENIA, plant-level)
  - ▶ Fabricated metal products (ISIC 381)
  - ▶ years 1979–1996
  - ▶ plant-year obs.  $\approx 4,000$
  - ▶ Materials: plant-reported intermediate consumption (*high precision*)
- ▶ U.S. (Compustat, firm-level)
  - ▶ Manufacturing (mostly durables incl. metal products; NAICS 33)
  - ▶ years 1970–2010
  - ▶ firm-year obs.  $\approx 8,000$
  - ▶ Materials: proxy via COGS – XLR; mixes labor/overhead (*lower precision*)
- ▶ Model & estimators.
  - ▶ Cobb–Douglas in  $k, l, m$
  - ▶ Control function: third-order polynomial in  $k, l, m$
  - ▶ Estimation: ACF baseline + CUE

# Elasticities & bootstrap normality — Chile (ISIC 381)

Table 3: Production function estimates (Chile, ISIC 381)

Parameter	Estimates		95% CI		BS Normality <sup>3</sup>
	ACF	CUE	Nonrobust <sup>1</sup>	Robust <sup>2</sup>	$p$
$\beta_k$	0.064	0.047	[0.014, 0.080]	[0.003, 0.087]	0.0
$\beta_l$	0.122	0.053	[-0.090, 0.195]	[-0.083, 0.185]	0.0
$\beta_m$	0.875	0.956	[0.834, 1.078]	[0.848, 1.072]	0.0
Returns to scale	1.060	1.060	—	—	—

<sup>1</sup> Wald. <sup>2</sup> Subset  $S$  Stock and Wright (2000). <sup>3</sup> Shapiro–Wilk (not reported).

- ▶ ACF and CUE estimates in ballpark
- ▶ Non-robust and robust CI similar and rel. tight
- ▶ Normality is rejected still

# Elasticities & bootstrap normality — U.S. (NAICS 33)

Table 4: Production function estimates (U.S., NAICS 33)

Parameter	Estimates		95% CI		BS Normality <sup>3</sup>
	ACF	CUE	Nonrobust <sup>1</sup>	Robust <sup>2</sup>	$p$
$\beta_k$	0.422	0.187	[-0.660, 1.034]	[-0.153, 0.507]	0.0
$\beta_l$	0.411	0.333	[-1.135, 1.801]	$(-\infty, 0.787]$	0.0
$\beta_m$	0.237	0.219	[0.106, 0.332]	[0.052, 0.332]	0.0
Returns to scale	1.070	0.739	—	—	—

<sup>1</sup> Wald. <sup>2</sup> Subset  $S$  Stock and Wright (2000). <sup>3</sup> Shapiro–Wilk.

- ▶ CUE aligns less with ACF
- ▶ robust CIs are large, unbounded for  $\beta_l$  (weak ID)
- ▶ Normality again rejected

# Conclusion

## Findings

- ▶ Proxies with poor explanatory power (weak proxies) are source of weak identification
- ▶ Monte Carlo shows potential large imprecision and non-normality in case of weak proxies
- ▶ Robust methods ensure valid inference

## Outlook

- ▶ Pretest that does not rely on BS (Andrews, 2018)
- ▶ More powerful robust inference (Kleibergen, 2005; Moreira, 2003)
- ▶ Other empirical applications

## References

- Akerberg, D. A., Caves, K., and Frazer, G. (2015). Identification properties of recent production function estimators. *Econometrica*, 83(6):2411–2451.
- Andrews, I. (2018). Valid two-step identification-robust confidence sets for gmm. *The Review of Economics and Statistics*, 100(2):337–348.
- Angelini, G., Cavaliere, G., and Fanelli, L. (2024). An identification and testing strategy for proxy-svars with weak proxies. *Journal of Econometrics*, 238(2):105604.
- Autor, D., Dorn, D., Katz, L. F., Patterson, C., and Van Reenen, J. (2020). The Fall of the Labor Share and the Rise of Superstar Firms. *The Quarterly Journal of Economics*, 135(2):645–709.
- De Loecker, J., Eeckhout, J., and Unger, G. (2020). The rise of market power and the macroeconomic implications. *The Quarterly Journal of Economics*, 135(2).
- Doornik, J. A. and Hansen, H. (2008). An omnibus test for univariate and multivariate normality. *Oxford Bulletin of Economics and Statistics*, 70(5):927–939.
- Gandhi, A., Navarro, S., and Rivers, D. A. (2020). On the identification of gross output production functions. *Journal of Political Economy*, 128(8). Electronically published.
- Gopinath, G., Kalemli-Özcan, , Karabarbounis, L., and Villegas-Sanchez, C. (2017). Capital Allocation and Productivity in South Europe. *The Quarterly Journal of Economics*, 132(4):1915–1967.
- Hansen, L. P., Heaton, J., and Yaron, A. (1996). Finite-sample properties of some alternative